Semi-streaming Quantizaton for Remote-sensing Data

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Abstract

We describe a strategy for reducing the size and complexity of very large, remote sensing data sets acquired from NASA's Earth Observing System. We apply the quantization paradigm from, and algorithms developed in signal processing, to the problem of summarization. Since data arrive in discrete chunks, we formulate a semi-streaming strategy that partially processes chunks as they become available, and stores the results. At the end of the summary time period, we reingest the partial summaries, and summarize them. We show that mean squared errors between the final summaries and the original data can be computed from the mean squared errors incurred at the two stages without directly accessing the original data. The procedure is demonstrated using data from JPL's Atmospheric Infrared Sounder.

1 Introduction

The motivation for this work is the need to facilitate exploratory data analysis of very large data sets produced by NASA Earth Observing System (EOS) satellites. EOS intends these data to be used by the greater research community in the study of Earth's climate system. However, for many researchers these data are too voluminous for the type of interactive, exploratory data analysis needed to formulate hypotheses and point the way to more detailed investigations. To ameliorate this problem, NASA instrument teams produce low volume, lower resolution, summary data sets typically comprised of means, standard deviations and other simple statistics for certain variables over an appropriate time period, and at coarse spatial resolution. For example, typical summary products might be aggregated daily or monthly over half, one, or five degree latitude-longitude spatial grid cells. Data processing constraints make this strategy

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attractive because means and standard deviations can be calculated in a "streaming" mode: one simply accumulates totals for each grid cell over the summary time period, and performs the appropriate division. Unfortunately, almost all distributional information is lost in this process, and this may be the information of greatest interest. For instance, subtle changes with time and space in the number of modes or occurrences of outliers may reflect important physical changes.

In Braverman (2002) we proposed a method for producing nonparametric, multivariate distribution estimates by grid cell. The method is an adaptation of the Entropy-constrained Vector Quantization (ECVQ) algorithm developed for signal processing by Chou, Lookabaugh and Gray (1989). The adaptation in Braverman (2002) was to make the method practical for large geophysical data sets and our data processing constraints. We demonstrated the method on a month's data from a fairly small but important cloud data set for half the northern hemisphere. Data from two EOS satellites, Terra and Aqua, have come online, and we now have to deal with more challenging circumstances.

In this paper we propose and demonstrate a semi-streaming version of the ECVQ-based method using three days of global data from a JPL instrument on Aqua, the Atmospheric Infrared Sounder (AIRS). Whereas the original method in Braverman (2002) required ingesting all data be summarized at once, the new method processes the data in chunks. A chunk is some amount of data which can be held in memory. Each chunk is partially summarized, and set aside. After all chunks are processed, the results are combined and summarized again. For example, in this paper we use a three day chunk of AIRS data which is partially summarized by applying a randomized version of the K-means algorithm to data in each grid cell. The cluster centroids, numbers of members, and within-cluster mean squared errors are recorded, and constitute a representation of the original data. These are then further summarized by applying a randomized version of ECVQ modified to account for input data with different masses. The example here uses a single three-day chunk, but in practice we envision summarizing a month of data in ten, three day chunks at the first stage. The advantage of this strategy is ingestion data in parts, so ECVQ can then be applied to partially reduced data. ECVQ requires some experimentation to set algorithm parameters, an unwieldy process if done on more voluminous, original data. The disadvantage is that the final summaries' mean squared errors may be higher. The consequences depend on the analysis for which the summary data will be used. In our example below, results appear acceptably close to those obtained had the analysis been performed on the original data.

This paper is organized as follows. In Section 2 we introduce the AIRS instrument and data. In Section 3 we discuss relevant aspects of quantization theory and how they apply to our statistical model of the relationship between original and summarized, or compressed, data. We also show how the practical constraints of our problem lead to the two-stage, semi-

streaming approach. In Section 4 we derive an expression for the mean squared error of the final summarized data relative to the original from the mean squared errors in the two processing stages. In Section 5 we report a sample data analysis using summarized AIRS data, and in Section 6 we discuss our conclusions.

2 Atmospheric Infrared Sounder Data and Science

Data used in this exercise are from the Atmospheric Infrared Sounder (AIRS). AIRS is a JPL instrument on-board NASA's EOS-Aqua satellite, launched on May 4, 2002. Aqua is in sun synchronous, polar orbit 705 kilometers above Earth, and crosses the equator during the ascending, or northward, part of its orbit at 1:30 pm local time. AIRS successively scans across a 1500 kilometer field of view taking data in 90 circular footprints as shown in Figure 1. As the spacecraft advances, the sensor resets and obtains another scan line. 135 scans are completed in six minutes, and this 90 by 135 footprint spatial array constitutes a "granule" of data. AIRS collects 240 granules per day, 120 on the day-time, ascending portions of orbits and 120 on night-time, descending portions. Granule ground footprints precess so granule 1 on a given day is not coincident with granule 1 on the next day. The descending granule map for July 20, 2002 is shown in Figure 2 as an example. At each of the 12,150 footprints in a granule, AIRS observes Earth and its atmosphere in 2378 infrared spectral channels. Roughly speaking, the channels sense the surface or different altitudes in the atmosphere. The instrument counts photons at the different wavenumbers, or inverse wavelengths. These counts are converted to brightness temperatures ranging from zero to about 340 degrees Kelvin. Certain atmospheric characteristics are related to photon emission, and these characteristics can be retrieved by solving complex sets of equations.

In this paper we use AIRS data acquired July 20 through 22, 2002 for channels listed in Table 1. The channels sample much of the vertical range of the atmosphere. The test data set contains the following variables: latitude (u), longitude (v), and eleven measurements at the wavenumbers shown in Table 1 $\boldsymbol{x}'_n = \{x_{n1}, x_{n2}, \ldots, x_{n11}\}$ for descending granules only. Each observation represents one AIRS footprint on the ground, and the test data comprise about 850 MB. The science question we address is whether the variance among channels 2 through 11 provides information about the presence of clouds. As AIRS looks down at Earth, different channels observe different levels of the atmosphere. If a scene is cloud-free, variation in brightness temperature across channels should be evident. If clouds are present, the view is interrupted and all channels see the same thing, namely the cloud. In this case, the variance should be lower. Channel 1 is excluded because it sees an atmospheric level above that of clouds.

Channel	*Wavenumber (cm$^-$1)*	*Sensitivity*
1	724.742	CO_2, upper troposphere
2	735.607	CO_2, upper troposphere
3	755.237	CO_2, upper troposphere
4	917.209	surface
5	1231.190	H_2O, deep troposphere
6	1285.323	H_2O, upper troposphere
7	1345.174	H_2O, upper troposphere
8	2412.562	CO_2, low troposphere
9	2450.020	surface
10	2500.313	surface
11	2616.095	surface

Table 1: AIRS channels included in test data, and the geophysical characteristics to which they are sensitive. Wavenumbers are inverse wavelengths.

To summarize the test data, we partition them according to membership in $1^\circ \times 1^\circ$ spatial grid cells, and reindex the $\boldsymbol{x}_n$'s as $\boldsymbol{x}_{uvn}$ where $u = -90, -89, \ldots, +89$ indexes latitude, $v = -180, -179, \ldots, +179$ indexes longitude, and n indexes 11-dimensional observations within grid cell. Let $N_{uv\cdot}$ be the total number of observations in the grid cell located at (u, v). At one degree resolution average channel variation using channels 2 through 11 in grid cell (u, v) is

$$w_{uv\cdot} = \frac{1}{N_{uv\cdot}} \sum_{n=1}^{N_{uv\cdot j}} \frac{1}{10} \sum_{i=2}^{11} (x_{uvni} - \bar{x}_{uvn\cdot})^2, \tag{1}$$

where i indexes channel components of the observation vector $\boldsymbol{x}_{uvni}$, and a "$\cdot$" indicates that a quantity is computed over all values of the corresponding index. While in principle it is possible to calculate quantities like $w_{uv\cdot}$ exactly, in practice it can be difficult to do so in the context of real-time, interactive exploratory data analysis. For example, if one wants to see what happens if channel 1 is included, the entire analysis has to be redone from scratch.

Braverman (2002) suggested and demonstrated a method for producing reduced volume and complexity proxy data sets that could be used in place of the original data for interactive, exploratory data analysis. The basic idea is as follows. We first spatially stratify the data by partitioning them into disjoint subsets based on membership in 1° spatial grid cells. Then, we replace the data in each cell with a smaller collection of representative, differentially weighted vectors. For example, if a grid cell subset has 1000 observations, we may replace it with five representative data points having weights 300, 300, 200, 100 and 100 respectively. The first

representative stands in for 300 of the original data values, the second for 300, the third for 200, and so on. The representatives vectors are centroids of points they represent. The set of observations sharing the same representative form a cluster. We call the set of representatives, their weights, and within cluster mean squared errors a "summary" of the subset. The collection of summaries for all $180 \times 360 = 64,800$ subsets is a compressed, or quantized, version of the data.

Figures 3 and 4 describe our new algorithm for creating grid cell summaries. This version is easier to implement within the constraints of our data processing system than is the algorithm described in Braverman (2002) because this version is semi-streaming: we process the data in chunks as they arrive, and set the results aside. Then, we read these "pseudodata" back in, and summarize them in a second stage of processing. This avoids the need to hold all data in memory at once. Our data processing requirements make it impossible to accumulate an entire month of data into memory. Instead, we accumulate data collected during the first three days of the month, process them, and store the results. This repeated for each three day period, called a "chunk". At the end of the month we read the chunk summaries back in as a proxy for the original data, and summarize them. In a month with 28, 29, or 31 days, the last chunk may be larger or smaller than the others. The monthly summaries generated this way may not be as accurate as those generated using a single stage implementation, but this is the price of applying the method in practice.

A key requirement of the two-stage method is that we be able to calculate the mean squared error between the final summaries and the original data without going back to the original data. As shown in Section 3.1, this is indeed possible by combining the mean squared errors from the two stages. First, however, we establish a statistical framework for describing, evaluating and comparing summaries.

3 Quantization

Our methodology borrows from quantization and souce coding theory. In this section we review relevant background material, and discuss a statistical model describing the relationship between original and summarized data. We then introduce practical considerations and modifications necessary to accommodate them. In particular, the framework introduced in this section explicitly addresses the effects of limiting output file sizes, and the need for summaries to be comparable to one another with respect to distributional features.

From here on, where it is clear we are discussing a single, one degree data subset, we drop the u and v spatial indices from the notation.

3.1 A Statistical Model for Quantized Data

Consider data belonging to a single one degree grid cell. In the case of the AIRS test data, this is a collection of N, eleven-dimensional data points, $\boldsymbol{x}_1, \boldsymbol{x}_2, \ldots, \boldsymbol{x}_N$. $\boldsymbol{x}_n = (x_{n1}, x_{n2}, \ldots, x_{n11})'$, where prime indicates vector transpose. The empirical distribution function of the data assigns probability $1/N$ to each $\boldsymbol{x}_n$, and for conceptual purposes we let $\boldsymbol{X}$ be an eleven dimensional random vector having this distribution. To summarize these data, we group the $\boldsymbol{x}_n$'s into clusters and use the set of cluster mean vectors and their associated proportions of members as a coarsened version of the original, empirical distribution. Figure 5 illustrates the basic idea.

Suppose $\boldsymbol{Y}$ is a random vector possessing the coarsened distribution. Then $\boldsymbol{Y}$ is a function of $\boldsymbol{X}$ determined by the way in which the $\boldsymbol{x}_n$'s are grouped. In other words, the support of $\boldsymbol{Y}$'s distribution is $\{\boldsymbol{y}_1, \boldsymbol{y}_2, \ldots, \boldsymbol{y}_K\}$, $K < N$, with $\boldsymbol{y}_k$ the average of all $\boldsymbol{x}_n$ belonging to cluster k:

$$\boldsymbol{y}_k = \frac{1}{M_k} \sum_{n=1}^{M_k} \boldsymbol{x}_n 1[\alpha(\boldsymbol{x}_n) = k].$$

$\alpha(\boldsymbol{x}_n)$ is an integer-valued function identifying to which of the K clusters $\boldsymbol{x}_n$ is assigned: if $\boldsymbol{x}_n$ is assigned to cluster k, then $\alpha(\boldsymbol{x}_n) = k$. $1[\cdot]$ is the indicator function, and $M_k = \sum_{n=1}^{N} 1[\alpha(\boldsymbol{x}_n) = k]$. The mass associated with $\boldsymbol{y}_k$ is M_k/N. In terms of random vectors, $\boldsymbol{Y} = E(\boldsymbol{X}|\alpha(\boldsymbol{X})) = E(\boldsymbol{X}|\boldsymbol{Y})$. Tarpey and Flury call this condition self-consistency of $\boldsymbol{Y}$ for $\boldsymbol{X}$ (Tarpey and Flury, 1996), and it implies that $\boldsymbol{y}_k$ is a minimum mean squared error estimate of all $\boldsymbol{x}_n$'s belonging to cluster k regardless of how the $\boldsymbol{x}_n$'s are assigned to clusters.

The accuracy with which $\boldsymbol{Y}$ represents $\boldsymbol{X}$ is measured by mean squared error (MSE), also called distortion δ:

$$\delta(\boldsymbol{Y}, \boldsymbol{X}) = \frac{1}{N} \sum_{n=1}^{N} \|\boldsymbol{x}_n - \boldsymbol{y}_{\boldsymbol{\alpha}}(\boldsymbol{x}_n)\|^2 = E\|\boldsymbol{X} - \boldsymbol{Y}\|^2 = trCov(\boldsymbol{Y} - \boldsymbol{X}),$$

where $\boldsymbol{y}_\alpha(\cdot)$ is the representative of the cluster to which $\boldsymbol{x}_n$ is assigned by α, and $\|\cdot\|$ is the usual vector norm. δ thus provides an upper bound on all elements of the error covariance matrix, $Cov(\boldsymbol{Y} - \boldsymbol{X})$, and this may be useful in propagating uncertainties through calculations that use $\boldsymbol{Y}$ as a proxy for $\boldsymbol{X}$. Moreover,

$$trCov(\boldsymbol{Y} - \boldsymbol{X}) = trE(\boldsymbol{X}\boldsymbol{X}') - 2trE(\boldsymbol{X}\boldsymbol{Y}') + trE(\boldsymbol{Y}\boldsymbol{Y}') = trE(\boldsymbol{X}\boldsymbol{X}') - trE(\boldsymbol{Y}\boldsymbol{Y}'),$$

since $E(\boldsymbol{X}\boldsymbol{Y}') = E[E(\boldsymbol{X}\boldsymbol{Y}'|\boldsymbol{Y})] = E[E(\boldsymbol{X}|\boldsymbol{Y})\boldsymbol{Y}'] = E(\boldsymbol{Y}\boldsymbol{Y}')$. This implies that the minimum MSE assignment of $\boldsymbol{x}$'s to clusters is the one that maximimzes $trE(\boldsymbol{X}\boldsymbol{Y}')$, which in turn maximizes $trCov(\boldsymbol{X}, \boldsymbol{Y}) + c$ with $c = E(\boldsymbol{Y}'\boldsymbol{Y}) = E(\boldsymbol{X}'\boldsymbol{X})$. K-means and similar algorithms iteratively search for such clusterings.

We measure data reduction in terms of the entropy of the distribution of $\boldsymbol{Y}$:

$$h(\boldsymbol{Y}) = -\sum_{k=1}^{K} P(\boldsymbol{Y} = \boldsymbol{y}_k) \log P(\boldsymbol{Y} = \boldsymbol{y}_k) = -\sum_{k=1}^{K} \frac{M_k}{N} \log \frac{M_k}{N} = -E \log p(\boldsymbol{Y}).$$

where $p(\boldsymbol{Y})$ is a function of the random vector $\boldsymbol{Y}$ realizing the value $P(\boldsymbol{Y} = \boldsymbol{y})$ when $\boldsymbol{Y} = \boldsymbol{y}$. A natural measure of data reduction is the difference in entropies of the distributions of $\boldsymbol{X}$ and $\boldsymbol{Y}$, but since $h(\boldsymbol{X}) = \log N$ is fixed, data reduction is maximized when $h(\boldsymbol{Y})$ is minimized. Entropy is a well recognized measure of information content (Ash, 1969, for example), and so we are quantifying data reduction as a decrease information.

Reducing information content comes at the cost of higher MSE (Cover and Thomas, 1991). Our goal is to find an assignment function, α, that simplifies the representation of the data by reducing $h(\boldsymbol{Y})$ relative to $h(\boldsymbol{X})$, and but does not incur large $\delta(\boldsymbol{Y}, \boldsymbol{X})$. Source coding theory, discussed in the next section, provides a formal framework for studying this trade-off.

3.2 Rate-distortion Characteristics

Source coding theory in signal processing provides a formal framework for studying the balance between quality of representation and data reduction. Source coding theory deals with the following problem. A stream of stochastic signals, $\boldsymbol{X}_1, \boldsymbol{X}_2, \ldots$ is to be sent over a channel for which the average number of bits per transmission is limited. It is impossible to transmit the signals exactly, since this would require bit strings longer than allowed by the capacity. So signals are encoded by assigning them to groups, and group indices sent instead. An encoding function, $\alpha(\boldsymbol{X})$, returns an integer in $\{0, 1, \ldots, K-1\}$ specifying group membership. This requires no more than $\log K$ bits per transmission where the logarithm is base 2. If the groups are labeled in a way that assigns low index numbers to the most frequently occurring groups, then the average number of bits per transmission is minimized and equal to the entropy of the distribution of the indices, $h(\alpha(\boldsymbol{X}))$. At the receiver, indices are replaced by group representatives as estimates of the original signals. The minimum mean squared error representatives are the group means, $\beta(k)$. Figure 6 is a schematic diagram of the procedure. The MSE of the estimate is $\delta(\boldsymbol{X}, \beta(\alpha(\boldsymbol{X})) = E\|\boldsymbol{X} - \beta(\alpha(\boldsymbol{X}))\|^2$. Assuming bits are assigned to groups in a way that minimizes entropy and minimum mean squared error representatives are used, the characteristics of the system are determined by the distribution of the signals and the encoder, α.

Every signal stream, or information source in the language of signal processing, has an associated function that describes the best combinations of distortion and entropy that can be obtained for it. The distortion-rate function is

$$\delta(h_0) = \min_{\alpha: h(\boldsymbol{Y}) \leq h_0} \delta(\boldsymbol{X}, \boldsymbol{Y}). \tag{2}$$

This function is non-increasing (Cover and Thomas, 1991), which conforms to our intuition that the more complex a clustering is, the better will be the quality of the representation. The function is also convex: it becomes progressively more costly to reduce error as entropy

increases. $\delta(h_0)$ provides the distortion-entropy combinations that are theoretically possible for signals from a given source distribution.

In this application we identify the signal stream with independent, random draws from a data set, and α with a clustering of its data points. We never actually draw from the data set or transmit a signal, but rather use the distortion-rate paradigm as a way to quantify the choices available in clustering the data. Two issues must be overcome in order to make use of the distortion-rate paradigm in this way: we must estimate the distortion-rate function, and determine which among the possible distortion-entropy combinations is the best for our purposes. It's easiest to discuss the latter issue first, as if the theoretical distortion-rate function were known, then to return to the problem of estimating $\delta(h_0)$.

The left panel of Figure 7 is a graphical device for comparing clusterings of a data set. We can represent any clustering as a point in the (δ, h) plane by calculating its $\delta(\boldsymbol{X}, \boldsymbol{Y})$ and $h(\boldsymbol{Y})$ values. If α_1 and α_2 are two such clusterings, α_1 is unambiguously better than α_2 if $\delta_1 < \delta_2$ and $h_1 \leq h_2$, or $h_1 < h_2$ and $\delta_1 \leq \delta_2$. α_2 is unambiguously better if $\delta_2 < \delta_1$ and $h_2 \leq h_1$, or $h_2 < h_1$ and $\delta_2 \leq \delta_1$. Otherwise, whether one is better than the other depends on whether one values low distortion more than low entropy. This is depicted by the two lines with different slopes, $-\lambda_1$ and $-\lambda_2$, which represent two different relative evaluations.

Relative importance of entropy and distortion can be specified as the increase in distortion we are willing to accept in exchange for a one bit reduction in the entropy of $\boldsymbol{Y}$: $\lambda = -\Delta\delta/\Delta h$. A line in (δ, h)-space with slope $-\lambda$ therefore expresses a set of equally desirable clusterings. The endpoints of the line give the maximum tolerable distortion or entropy, and points on the line represent all allowable combinations of the two. Such lines can be thought of as a sort of compression budget, describing the distortion-entropy combinations we are willing to accept. All lines with the same slope represent the same distortion-entropy trade-off, but lines shifted closer to the origin are better in the sense that they represent combinations with lower expenditures.

The right panel of Figure 7 shows a typical distortion-rate function when the number of clusters, K, is allowed to range between one and the number of data points, N. This case is labeled $\delta(h), K = N$. Two extreme results are possible. First, every data point is assigned to its own cluster in which case $\delta = 0$ and $h = \log N$. Second, all points are assigned to a single cluster in which case $\delta = N^{-1} \sum_{n=1}^{N} \|\boldsymbol{x}_n - \bar{\boldsymbol{x}}\|^2 = trCov(\boldsymbol{X})$. $\bar{\boldsymbol{x}}$ is the data set centroid. If the number of clusters is restricted to $K < N$, the distortion-rate function shifts in the direction of greater distortion because at least one cluster will have to have at least two points assigned to it. Note that $\delta(0)$ remains fixed because $h = 0$ implies all $\boldsymbol{x}$ are assigned to one cluster regardless of K. In the implementation discussed later it will be necessary to impose such a restriction to ensure that output files are within size limitations. In either case, the optimal position along

$\delta(h)$ is where a line with slope $-\lambda$ is tangent to $\delta(h)$, since this is the clustering at which the trade-off between distortion and entropy in the data is equal to the desired rate given by $-\lambda$.

We estimate the rate-distortion function for a data set by evaluating the function at discrete points. One way to do this is to choose a set of h values, and for each find the minimum distortion clustering among all those with entropies less than or equal to each value of h. Instead of minimizing distortion subject to a constraint on entropy, we use a randomized version of the ECVQ algorithm modified for large data sets (Braverman, 2002). We minimize the lagrangian loss function:

$$L(\lambda) = \sum_{n=1}^{N} \|\boldsymbol{x}_n - \boldsymbol{y}_\alpha(\boldsymbol{x}_n)\|^2 + \lambda \left[-\log \frac{M_\alpha(\boldsymbol{x}_n)}{N} \right], \tag{3}$$

where $M_\alpha(\boldsymbol{x}_n)$ is the number of data points assigned to the cluster to which $\boldsymbol{x}_n$ is assigned, and the logarithm is base two. The distortion-rate function is estimated by applying ECVQ to the data for a set of λ values. Of course, if we have already fixed λ *a priori*, there is no need to estimate $\delta(h)$ at other values of λ.

3.3 Practical Considerations

This discussion now comes full-circle because we admit we do not know λ *a priori*. In fact, without any further information the best choice of λ would be zero because this minimizes MSE. ECVQ then becomes equivalent to K-means. However, a unique, crucial aspect of this application creates a requirement that goes beyond the general need to achieve accuracy and parsimony. It provides an additional constraint that leads us to consider non-zero λ values.

The unique aspect here is that we are not summarizing a single data set in isolation; we are summarizing 64,800 grid cell data subsets in concert. These summaries need to be comparable to one another in the sense that conclusions drawn by comparing them should approximate the conclusions that would've been drawn by comparing the underlying, original grid cell data. Analyses based on grid cell summaries that appear different only because some provide better descriptions of their data than do others, are misleading. We must ensure that apparent differences among summaries are due to real differences in the data, not due to differences in how well summaries fit their data. This observation provides a criterion for setting λ in the loss function in Equation (3): choose λ to minimize the variance among distortions across grid cell subsets, and use the same λ in all subsets. By setting λ to equalize distortions as much as possible, we sacrifice quality in some summaries to achieve comparability across grid cells.

A significant drawback to this strategy is that λ must be found by experimentation. This requires repeated applications of the ECVQ algorithm in all, or at least a representative sample of grid cells as described in Braverman 2002. Although the testing procedure can be parallelized

by running different grid cells on different processors, it is still computationally intensive. To mitigate the problem we capitalize on another aspect of this applicaton that might at first seem like an obstacle: data volume and their arrival schedule.

Recall that AIRS data are collected and processed by granule, and that a given grid cell receives data contributions from multiple granules acquired on multiple days. Data in that cell from the same granule are relatively close spatially, and are acquired at nearly the same time. In general we expect them to be more correlated with each other than with data from other granules or days. It therefore seems reasonable to summarize the granule contributions in chunks defined in units of days, as they arrive. In fact, this is a practical necessity. There is too much data to accumulate, say, an entire month before starting processing. In addition, we are unable to choose λ by the criterion discussed above without access to the complete data, and so we have no choice but to use $\lambda = 0$ at this point. The result is less uniform quality across grid cells than we would like, but we can correct for that in the second stage. The stage one output are small enough to permit experiments with λ.

Two modifications are necessary for this approach. First, the pseudo-data points have unequal masses, and this must be taken into account in computing centroids and evaluating the ECVQ loss function in the second stage. This is not difficult; it just requires computing weighted averages. Second, we need to be able to reconstruct the MSE between the raw data and the final summaries without going back to the raw data. Fortunately, this is possible, as shown next.

4 Two-stage Quantization

Suppose now we have separate subsets which have been previously summarized, and we wish to combine them. Assume we have T summaries from T chunks in a one degree grid cell. Each chunk provides one subset. A random draw from the tth subset is represented by the random vector $\boldsymbol{X}_t$ having the empirical distribution of that subset. Let the number of data points in subset t be N_t, and let V be an integer valued random variable such that $P(V = t) = N_t / \sum_{s=1}^{T} N_s$ independent of the $\boldsymbol{X}_t$'s. Let $\boldsymbol{Y}_t = q(\boldsymbol{X}_t)$ be the representative of $\boldsymbol{X}_t$.

Next, define $\boldsymbol{Y}$ and $\boldsymbol{X}$ as follows:

$$\boldsymbol{Y} = \sum_{t=1}^{T} \boldsymbol{Y}_t 1[V = t],$$

$$\boldsymbol{X} = \sum_{t=1}^{T} \boldsymbol{X}_t 1[V = t].$$

Let $\delta(\boldsymbol{Y}_t, \boldsymbol{X}_t)$ be the distortion resulting from the summarization of $\boldsymbol{X}_t$, and let $h(\boldsymbol{Y}_t)$ be the

entropy. We would've liked to cluster the combined data represented by $\boldsymbol{X}$, but we have only the summarized data represented by the $\boldsymbol{Y}_t$'s, and the resulting $\delta(\boldsymbol{Y}_t, \boldsymbol{X}_t)$'s. We create $\boldsymbol{W}$ by summarizing the $\boldsymbol{Y}_t$'s, and therefore have $\delta(\boldsymbol{W}, \boldsymbol{Y})$ also. We would like to reconstruct the distortion $\delta(\boldsymbol{W}, \boldsymbol{X})$ from the information we have available. Note that

$$\begin{aligned}\delta(\boldsymbol{W}, \boldsymbol{X}) &= E\|\boldsymbol{W} - \boldsymbol{X}\|^2 = E\|(\boldsymbol{W} - \boldsymbol{Y}) + (\boldsymbol{Y} - \boldsymbol{X})\|^2 \\ &= E\|\boldsymbol{W} - \boldsymbol{Y}\|^2 + E\|\boldsymbol{Y} - \boldsymbol{X}\|^2 + 2E[\boldsymbol{W}'\boldsymbol{Y} - \boldsymbol{W}'\boldsymbol{X} - \boldsymbol{Y}'\boldsymbol{Y} + \boldsymbol{Y}'\boldsymbol{X}] \\ &= \delta(\boldsymbol{W}, \boldsymbol{Y}) + \delta(\boldsymbol{Y}, \boldsymbol{X}). \end{aligned} \tag{4}$$

The crossterm in (4) is zero because $E(\boldsymbol{W}'\boldsymbol{Y}) = E(\boldsymbol{Y}'\boldsymbol{Y})$ and $E(\boldsymbol{W}'\boldsymbol{X}) = E(\boldsymbol{Y}'\boldsymbol{X})$:

$$\begin{aligned}E\left(\boldsymbol{W}'\boldsymbol{Y}\right) =& E\left(E(\boldsymbol{Y}|\boldsymbol{W})'\boldsymbol{Y}\right) = E\left(E(\boldsymbol{Y}'\boldsymbol{Y}|\boldsymbol{W})\right) = E\left(\boldsymbol{Y}'\boldsymbol{Y}\right), \\ E\left(\boldsymbol{W}'\boldsymbol{X}\right) =& E\left(E(\boldsymbol{Y}|\boldsymbol{W})'\boldsymbol{X}\right) = E\left[E\left(E(\boldsymbol{Y}|\boldsymbol{W})\right)'\boldsymbol{X}|\boldsymbol{X}\right] \\ =& E\left(E(\boldsymbol{Y}'\boldsymbol{X}|\boldsymbol{X}, \boldsymbol{W})\right) = E\left(\boldsymbol{Y}'\boldsymbol{X}\right).\end{aligned}$$

Therefore, the distortion between the grand summary constructed the individual subset summaries, $\boldsymbol{W}$, and the original data, $\boldsymbol{X}$, can be reconstructed as the sum of the distortions of the two stages. Moreover,

$$\begin{aligned}\delta(\boldsymbol{Y}, \boldsymbol{X}) &= E||\sum_{t=1}^{T} \boldsymbol{Y}_t 1[V = t] - \sum_{t=1}^{T} \boldsymbol{X}_t 1[V = t]||^2 \\ &= \sum_{t=1}^{T} E(\boldsymbol{Y}_t - \boldsymbol{X}_t)'(\boldsymbol{Y}_t - \boldsymbol{X}_t)P(V = t) \\ &= \sum_{t=1}^{T} \delta(\boldsymbol{Y}_t, \boldsymbol{X}_t)P(V = t).\end{aligned}$$

Thus, the total distortion of the two stage procedure, $\delta(\boldsymbol{W}, \boldsymbol{X})$, is the distortion from the second stage, $\delta(\boldsymbol{W}, \boldsymbol{Y})$, plus a weighted average of the distortions from the first stage with weights given by $P(V = t) = N_t^{-1} \sum_{s=1}^{T} N_s$:

$$\delta(\boldsymbol{W}, \boldsymbol{X}) = \delta(\boldsymbol{W}, \boldsymbol{Y}) + \sum_{t=1}^{T} \delta(\boldsymbol{Y}_t, \boldsymbol{X}_t) \frac{N_t}{\sum_{s=1}^{T} N_s}. \tag{5}$$

5 Creating and Analyzing Summarized AIRS Data

5.1 Computations

Figures 3 and 4 depict the two-stage algorithm used to summarize the AIRS test data in our example. Briefly, the first stage uses a randomized version of K-means clustering to partially reduce the data. We run K-means S times on each grid cell using a different random sample on each trial as initial cluster seeds. We choose the minimum MSE summary to represent the

original data. For the second stage, we concatenate the first stage summaries for the grid cell, and summarize them using a randomized version of the ECVQ algorithm modified to account for the fact that the data points have different masses. We use weighted averages to calculate cluster representatives, distortions, and the objective function. As in the first stage, ECVQ is run S times with different random samples as cluster seeds. For each of the resulting sets of representatives, we calculate the MSE between the second stage input and the set. The minimum MSE set is selected as the final summary of the data. Additional details concerning the basic version of the ECVQ algorithm can be found in Braverman (2002) and Chou, Lookabaugh and Gray (1989).

For this exercise, we processed data from 360 descending granules from July 20 through 22, 2002, comprising about 850 MB. The first-stage, randomized K-means used $S = 30$ trials and $K = 20$ clusters for all grid cells. This yields $\delta(\boldsymbol{Y}_t, \boldsymbol{X}_t)$ for $t = 1$. Here, we process the test data in one chunk for demonstration purposes. The second-stage, randomized ECVQ used $S = 30$ trials, $K = 15$ as the maximum number of clusters allowed, and $\lambda = .06$. Second stage processing yields $\delta(\boldsymbol{W}, \boldsymbol{Y})$. Algorithm parameter settings were determined as follows. λ was selected by experimentation as described in Section 3.3 using the first-stage output. For a more algorithmic description of how λ is determined, see Braverman (2002). For both stages, S was chosen as small as possible, but still according to the basic rule-of-thumb for generating "large" samples. The values of K were chosen as small as possible given we wanted to have some confidence that a randomly selected set of K data points would be representative of the two-dimensional sub-space spanned by the two most important principle components of the original data. These account for about 90 percent of the data set's variance. In fact, we performed both the K-means and ECVQ clustering procedures on data projected into that subspace. Resulting cluster assignments were used to compute new representatives in the original 11-dimensional space.

For timing comparisons, we selected a small spatial region in the western Pacific, and also ran the single stage version of the algorithm (see Braverman 2002) on these data. The one-stage version is analogous to what is shown in Figure 4 except that the original data are input rather than the first stage output. For this we used $K = 15$ and $\lambda = .06$ *a priori*. Table 2 provides some statistics concerning the original data, and the two implementations on an average per-grid-cell basis. CPU time does not include time necessary to determine the correct value of λ in the second stage. That involves experimenting with various candidate values and running the algorithm repeatedly on a sample of grid cells in the one-stage implementation, or on all grid cells in the two-stage version. Typically, the task takes about fifteen times longer for the one-stage implementation than is shown in the table. However, the comparison is academic

Characteristic	*Original data*	*One-stage implementation*	*Two-stage implementation*	
			after first stage	*after second stage*
Records	134.25	11.11	19.5	9.06
CPU seconds	NA	2.95	4.95	6.63
MSE	0	27.93	30.23	43.26

Table 2: Performance statistics for single and two-stage implementations. All quantities are on an average-per-grid cell basis. Note that the CPU time for the two-stage implementation is broken into two parts. The 'after second stage' figure is total time, and includes the 'after first stage' figure. Also, CPU times do not include time necessary to determine the proper value of λ.

because the one-stage implementation is simply not viable in our processing environment.

The estimated average value of w by grid cell is shown in Figure 9. The calculation uses the summarized data:

$$\hat{w}_{uv\cdot} = \frac{1}{N_{uv\cdot}} \sum_{k=1}^{K_{uv}} M_{uvk} \frac{1}{10} \sum_{i=2}^{11} (y_{uvki} - \bar{y}_{uvk\cdot})^2, \tag{6}$$

where uv indexes spatial location, K_{uv} is the number of clusters, y_{uvki} is the ith component of the kth cluster representative, and M_{uvk} is the number of observations in cluster k. The true value for the grid cell at latitude u and longitude v is $w_{uv\cdot}$, given in Equation (1).

Figure 10 is a map of the MSE's between summaries and the original data by grid cell, $\delta(\boldsymbol{W}, \boldsymbol{X})$, calculated using Equation (5):

$$\delta_{uv}(\boldsymbol{W}, \boldsymbol{X}) = \delta_{uv}(\boldsymbol{W}, \boldsymbol{Y}) + \delta_{uv}(\boldsymbol{Y}_1, \boldsymbol{X}_1).$$

Note that MSE is relatively uniform across most grid cells. The high MSE areas in Anarctica, the Himalayas, and along the equator appear due to zero values on channels 10 and 11 in the raw data; a known problem discussed below.

To get a feel for the significance of elevated mean squared errors in the equitorial region, we calculated the true values of w using Equation (1) for the grid cells in the study area identified by the white box in Figure 10. Figure 11 is a map of this area, which extends from 15° to 30° north latitude, and from 135° to 165° west longitude, showing the ratio $(\hat{w}/w)$. All values are less than or equal to one, consistent with Jensen's Inequality for convex functions. Even in this high MSE area, the ratio is near one for all but a handful of cells. This provides some measure of assurance that the information in Figure 9 is reliable.

5.2 Data Analysis: Characterizations of Cloudiness

We now ask whether the distributional characteristics expressed by grid cell summaries reveal meaningful information about the nature of physical phenomena that would otherwise be difficult to obtain. Figure 11 concentrates on the western Pacific in part because of the prominent feature in that area in Figure 9. The feature looks roughly like a backwards "c" of low $\hat{w}$ values embedded in an area of high values. We find a number of distributionally interesting grid cell summaries in that region. In particular, summaries for (16N, 158E) and (21N, 154E) are shown in Figure 12. The left panels show the cluster representatives in the form of parallel coordinate plots, and the right panels show numbers of cluster members in the form of histograms. In the histograms, cluster indices between 0 to $K_{uv} - 1$ are shown on the extreme right, and the count associated with each cluster is in green on the left. The colors of the bars match those of the lines of corresponding cluster representatives in the parallel coordinate plots. This is somewhat difficult to see in these static graphics. In our analysis we used an interactive tool to click on a line on the left or a bar on the right, and see the corresponding bar or line highlighted.

Figure 13 contains daily images of the AIRS test data at wavelength 980 cm^{-1} in the western Pacific study area. 980 cm^{-1} is relatively close to our channel 4, and is sensitive to either surface temperature, or to the temperature of any intervening clouds. A swirl of cold clouds with brightness temperatures less than 220 Kelvin is shown in blue in the July 21 image. It can also be seen on the right side of the July 20 image, and less clearly in the left side of the July 22 image. On July 21 a small warm spot exists in the center of the blue structure. This is the signature of a tropical storm or hurricane, including the warm, clear, eye. A check of tropical storm archives showed this to be Typhoon Fengshen, one of the most powerful storms of 2002. Fengshen moved slowly westward though our study area, and later caused heavy rainfall in Japan and Korea. A second, smaller tropical storm that did not attain typhoon status is also apparent to the west of Typhoon Fengshen on July 20.

The parallel coordinate plot for (16N, 158E) shows three types of clusters in 11-dimensional space. The top group generally exceeds 250 Kelvin, and shows the most variability across channels 2 through 11. Since they are warm and variable, this group corresponds to clear areas seen as yellow or red in Figure 13. The morphology of these parallel coordinate lines is typical of clear scenes in other areas where AIRS views to the ocean surface. Of the 75 data points in this grid cell, 22 are represented by these clear-scene clusters, and these observations appear to be contributed mostly on July 22. The bottom group of clusters in the parallel coordinate plot show fairly consistent temperatures on channels 2 through 11, and are very cold at about 200 Kelvin. This is the signature of high, cold clouds indicating severe storminess. These clusters represent 34 of 75 data points in (16N, 158E), and correspond the blue areas around that location in the

upper two panels of Figure 13. Finally, the middle group of clusters at about 225° indicate intermediate cases. Their parallel coordinate lines suggest cloudiness, but their temperatures suggest they are at middle or low altitudes. This is consistent with the aftermath of a storm, as indicated by the green areas in and around this grid cell on July 21 and 22.

The parallel coordinate plot for (21N, 154E) shows two types of cluster signatures, one with an unusual feature. The upper, warmer group of four clusters represent 9 of 49 data points in the grid cell, with 6 data points belonging to the warmest cluster. Three of these clusters exhibit intermediate variability across channels. The fourth contains one data point, distinguished by an unusual combination of low variability and warm temperature. Its signature crosses that of the adjacent cluster at channel 3, quite unusual among grid cell distributions. The data point is also warmer at upper altitudes compared to the other clusters, as shown by channels 1, 2, and 3, but cooler at lower altitudes, as shown by channel 8. Also, the underlying ocean surface is relatively cool, as shown by lower values on channels 9, 10, and 11. These thermal characteristics are consistent with the eye of Typhoon Fengsheng as seen in Figure 13.

The bottom set of clusters in Figure 13 clearly contain some anomalous data. Zero values on some channels are associated with poor signal to noise ratios, a common problem for short wavelength channels when viewing cold scenes. This is not expected to occur with warm scenes, and examination of the data summaries in other areas bear this out.

This analysis shows the scientific utility of summarizing AIRS data by quantization. The meteorological significance of these data is well understood, and the current challenge is to utilize the enormous amount of information they contain. Here we begin the process of understanding relationships between certain types of geophysical phenomena and distributional features of data generated by them.

6 Summary and Conclusions

In this paper we described and demonstrated a method for summarizing large remote-sensing data sets. It is a two-stage procedure in which the first stage is semi-streaming. That is, the data are processed in three-day blocks, called chunks. The first stage summaries are stored, and at the end of the summary period we summarize the summaries in the second stage of processing. We demonstrated the procedure with one chunk of data from NASA's AIRS instrument. We argued that a quantization method developed in signal processing applies as well to constructing data summaries. These summaries optimally balance reduction in complexity and size against increased error. In Section 3.1 we showed that the final summary errors can be determined from errors calculated in two stages, without direct access to the original data. This is a key feature

of streaming algorithms. Finally, we demonstrated the scientific utility of our method by using the summarized data to identify and study physical features of Typhoon Fengshen.

There are three conclusions from this exercise. First, the summaries do capture important distributional features of the data related to physical processes. We paid special attention to ensuring that summaries for different grid cells are comparable by carefully choosing the second-stage algorithm's λ parameter. A second conclusion is that additional modifications may be necessary to make this method practical for our data production environment. Five processing seconds per grid cell for the first stage translates to $((5 \times 64800)/3600 =)$ 90 hours per chunk. We can speed this up by lowering K, the number of clusters, lowering S the number of trials, or by using sampling in determining cluster representatives (Braverman, 2002). These strategies will most likely result in summaries with higher mean squared errors, although magnitudes are uncertain because they depend on data characteristics. We will perform experiments to determine how best to cope with this. Our third conclusion is that interactivity is key to making the best use of summarized data for exploratory science analysis. We made heavy use of a Java tool written especially for viewing summaries interactively. This allowed us to quickly compare different grid cells, understand how summaries change as a function of spatial location, and draw connections between physical phenomena and distributional characteristics. We are encouraged by the results of this exploration, and eagerly anticipate moving forward on this project.

Acknowledgements

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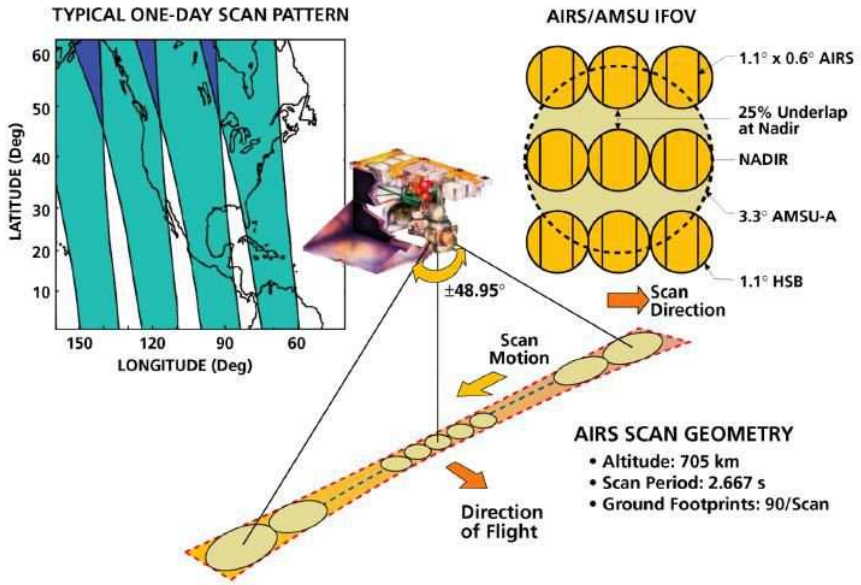

Figure 1: AIRS scan geometry. Left: Ground-tracks for ascending portion of orbit. Center: AIRS instrument and ground footprints in one scan line. Right: AIRS, AMSU and HSB views of one footprint. AMSU and HSB are two other instruments on Aqua.

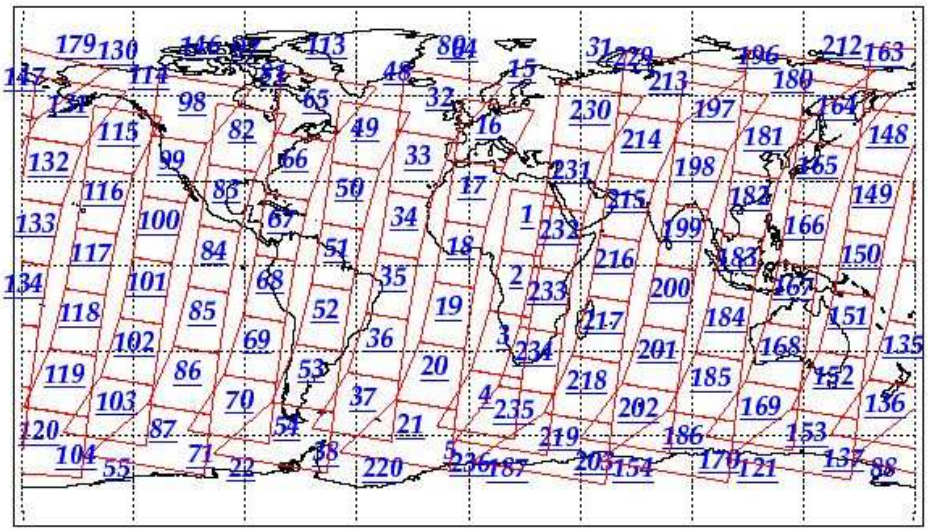

Figure 2: Descending granule map for July 20, 2002.

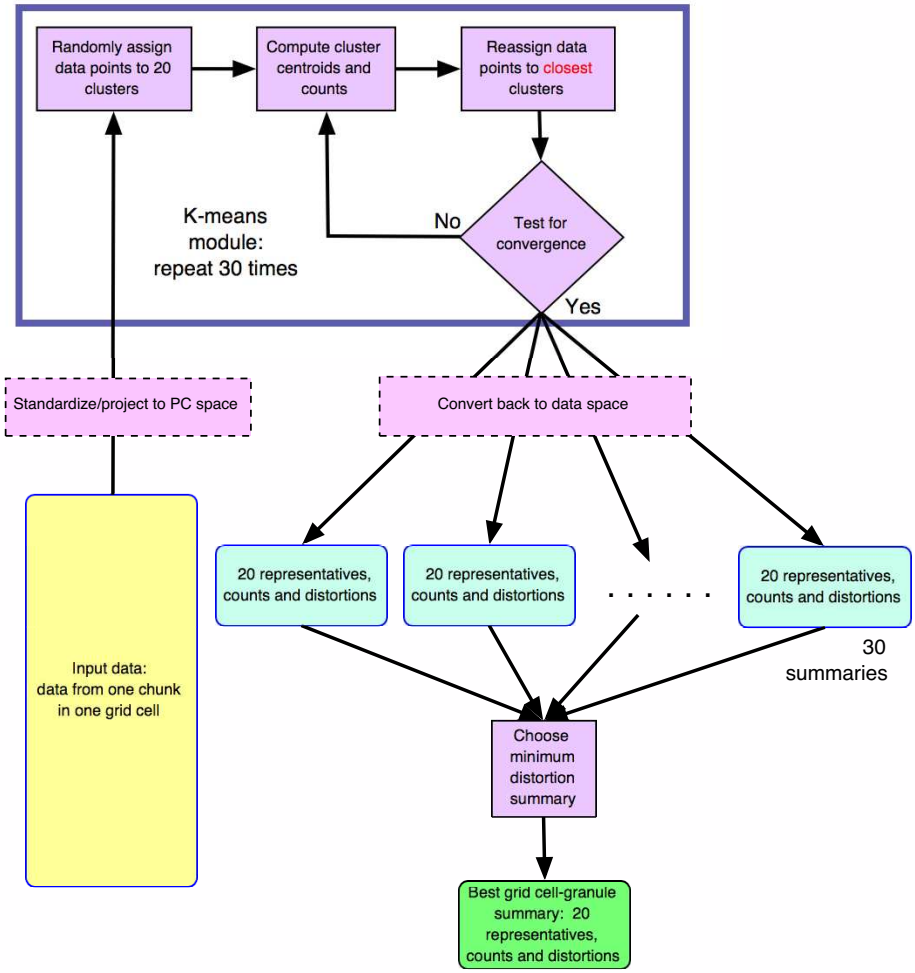

Figure 3: Schematic diagram depicting the first stage of quantization for a single 1° latitude by 1° longitude grid cell. Individual chunk subsets are clustered using a randomized version of the K-means algorithm. K-means is repeated $S = 30$ times, using a different set of randomly selected cluster seeds each time. The data summary is the set of K cluster centroids, and their associated numbers of members and within-cluster mean squared errors. The input data may be standardized or projected into a principal component subspace to reduce dimensionality. If so, the representatives are converted back to data space.

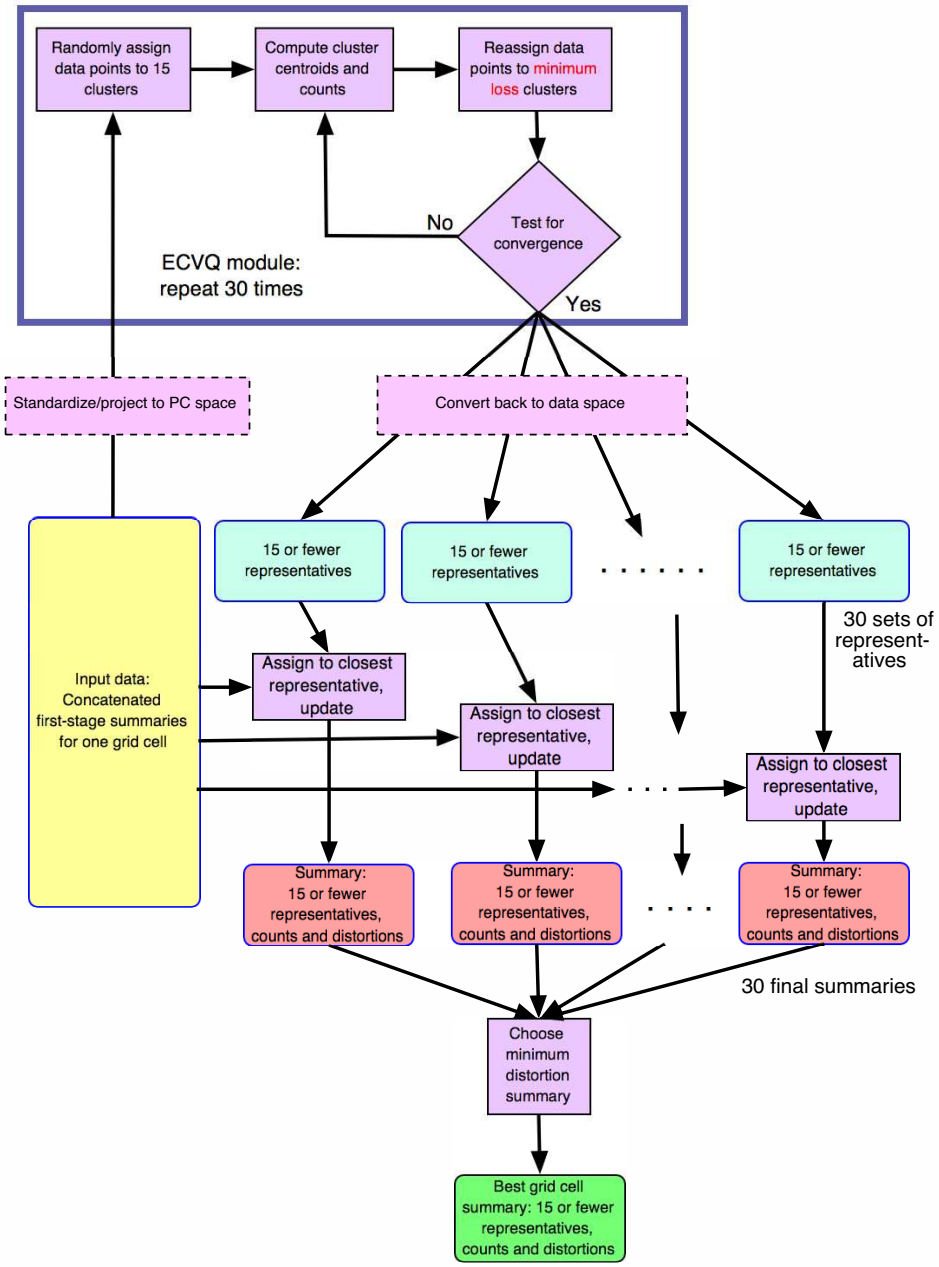

Figure 4: Schematic diagram depicting second stage quantization of single $1° \times 1°$ cell. The three individual first-stage summaries are concatenated to form a new input data set in which data points have different mass. These data are summarized using a randomized version of the ECVQ algorithm.

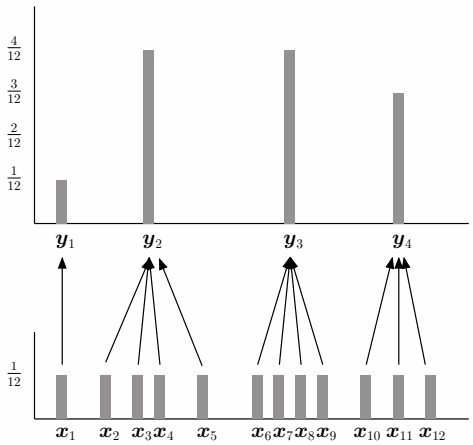

Figure 5: Schematic diagram of quantization. Although shown here as if scalar-valued, $\boldsymbol{x}$'s and $\boldsymbol{y}$'s are vectors.

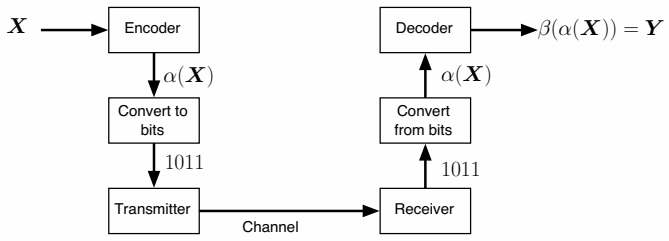

Figure 6: Schematic diagram of signal encoding, transmission and decoding.

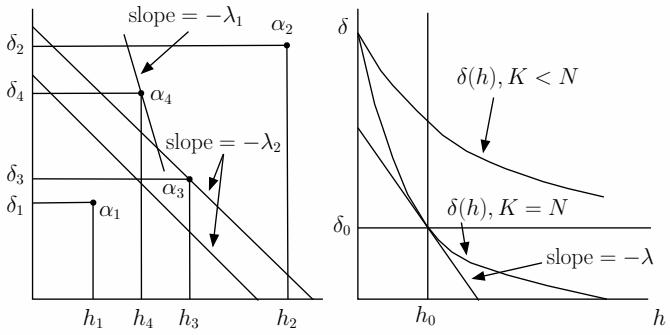

Figure 7: Left: rate-distortion plane. Entropy is on the horizontal axis and distortion is on the vertical axis. The clustering specified by α_1 is superior to the others because it has lower distortion and entropy. α_2 is inferior to the others because it has higher distortion and entropy. Which of α_3 and α_4 is better depends on whether low distortion is valued more than low entropy. If λ_1 reflects our relative valuation, then α_4 is preferred. If λ_2 reflects our relative valuation then α_3 is preferred. Note that all points along the line with slope $-\lambda_2$ are equally desirable, and lines parallel lines to the interior represent more desirable clusterings. Right: distortion-rate space showing $\delta(h)$, the distortion-rate function for a data set or information source. $\delta(h)$ shows the minimum possible distortion among all clusterings having entropy less than or equal to h. λ represents the rate at which one is willing to accept additional distortion for a reduction in entropy. The optimal location along $\delta(h)$ for a given λ is the tangent point of the line with slope $-\lambda$ with $\delta(h)$. Restricting the number of clusters to be less than the number of data points ($K < N$) causes $\delta(h)$ to shift in the direction of greater distortion.

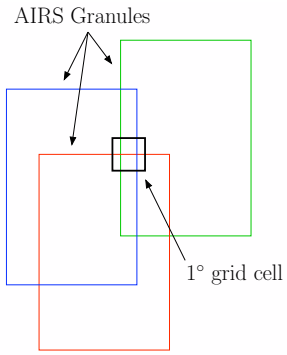

Figure 8: The spatial dimensions and relative locations of three chunks are shown in red, green and blue. The black square shows the spatial dimensions and relative location of a grid cell which receives data from all three. Here, the chunks are shown as different, overlapping granules as would be the case if chunks were separate days.

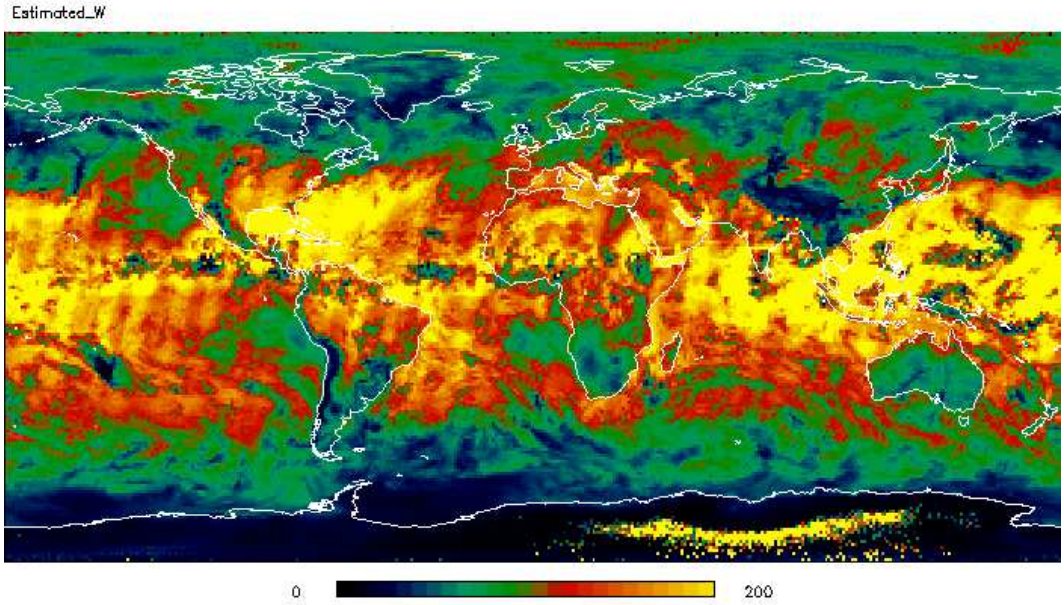

Figure 9: Estimated value of w from summarized data. Units are degrees Kelvin.

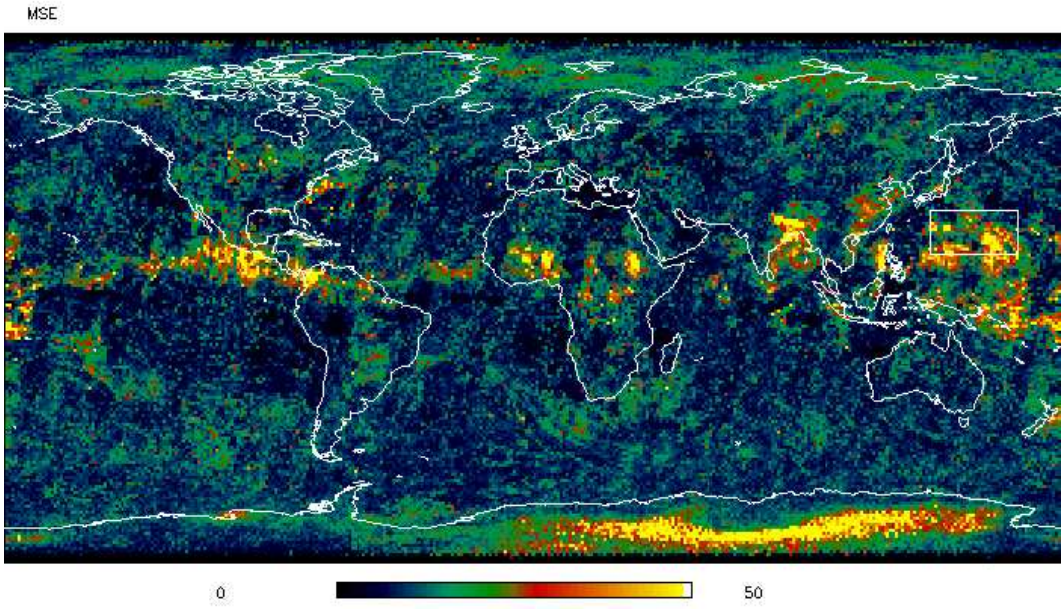

Figure 10: Mean squared error between summarized and original data. Units are degrees Kelvin. The white box shows the study area, 15N to 30N latitude, 135E to 165E longitude.

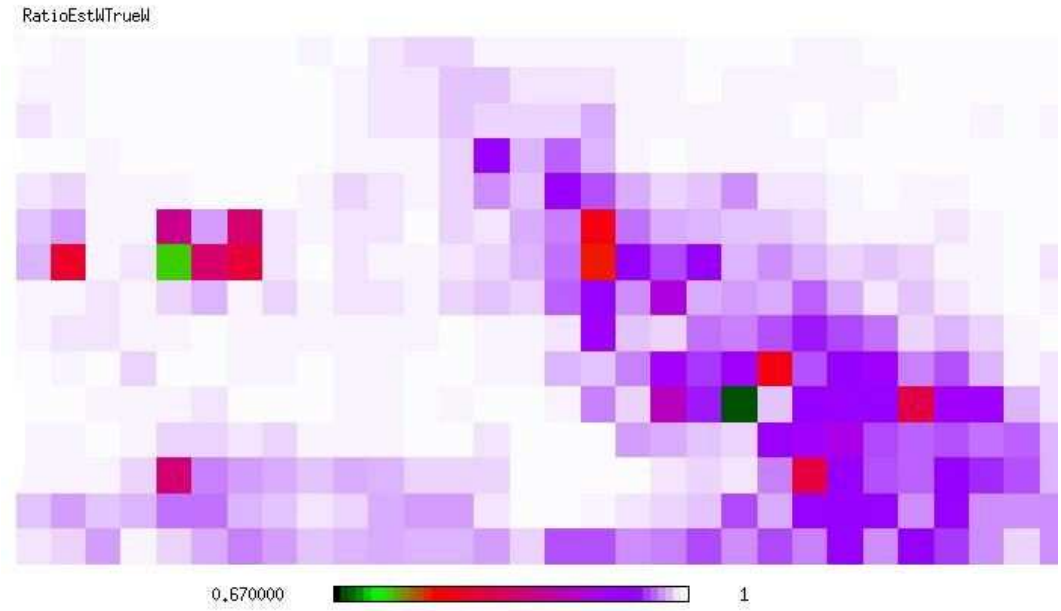

Figure 11: Ratio of estimated to true values of w ($\hat{w}/w$) for the western Pacific test area.

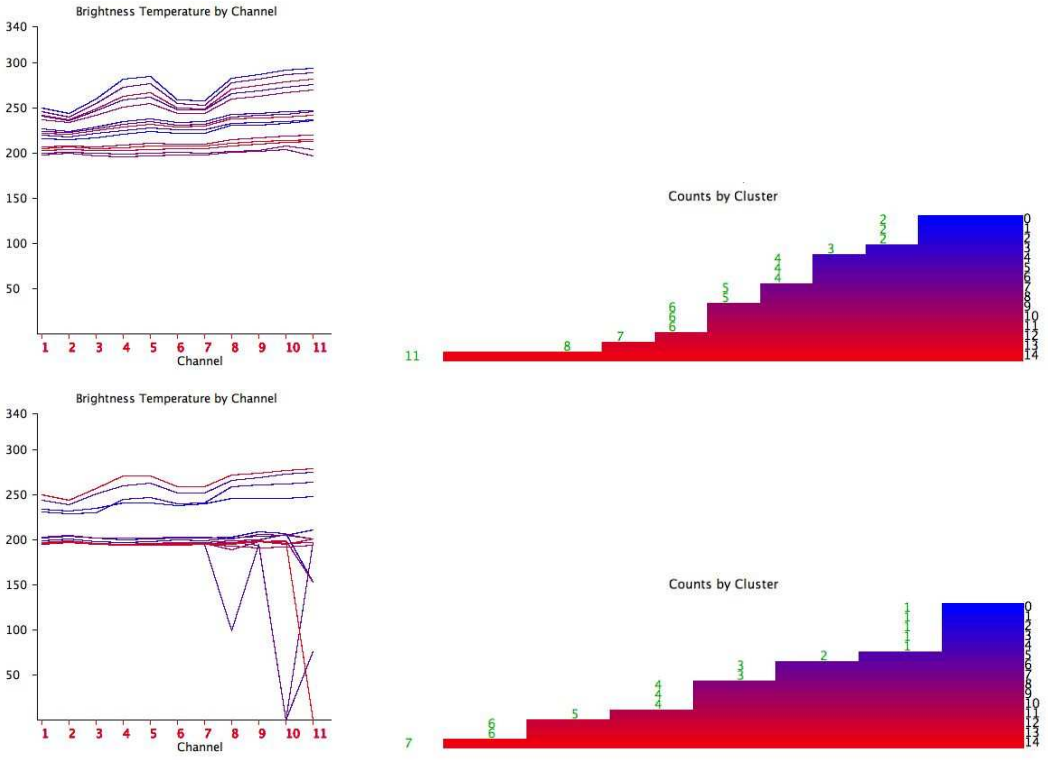

Figure 12: Cluster representatives, left, and counts, right for three grid cells. Top: 16N, 158E. Bottom: 21N, 154E.

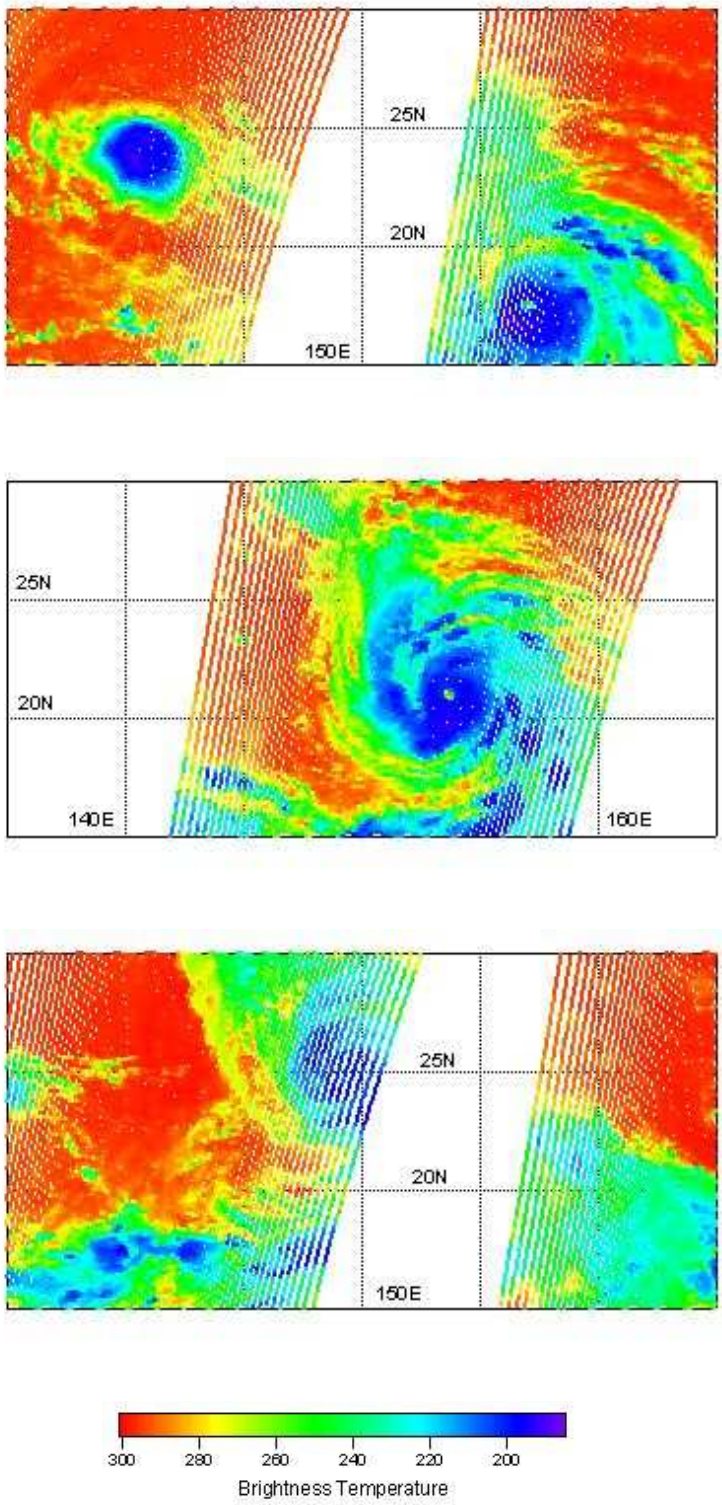

Figure 13: AIRS measurements at wavelength 980 cm^{-1} for the study region. Top: July 20. Middle: July 21. Bottom: July 22. Units are degrees Kelvin.